

Chernoff bounds and AMS

Last time:

We assumed access to sk 'rows' of M bits each, drawn uniformly at random from $\{-1, +1\}$.

$$\begin{array}{cccc} u_{11} & \dots & u_{1M} & \leftarrow \text{row} \\ u_{21} & \dots & u_{2M} & \\ \vdots & & & \\ u_{sk1} & \dots & u_{skM} & \end{array}$$

We showed that if $s \geq \frac{16}{\epsilon^2}$, then the estimator $Y_i = \left(\sum_{j=1}^M f_j u_{ij} \right)^2$ satisfies

$$\Pr [|Y_i - F_2| \geq \epsilon F_2] \leq \frac{1}{8}.$$

We now wish to choose k in such a way that the median of Y of Y_1, \dots, Y_k satisfies

$$\Pr [|Y - F_2| \geq \epsilon F_2] \leq \delta.$$

To do this, we defined X_i the random variables X_i :

$$X_i = \begin{cases} 1 & \text{if } Y \geq (1+\epsilon)F_2 \text{ or } |Y - F_2| \geq \epsilon F_2 \\ 0 & \text{otherwise.} \end{cases}$$

Now we have: \Pr

$$(1) \quad \Pr[X_i] \leq \frac{1}{8}$$

(2) X_i are iid

$$(3) \quad |Y - F_2| \geq \epsilon F_2 \Rightarrow \frac{1}{k} \sum_{i=1}^k X_i \geq \frac{1}{2}$$

Chernoff bounds

Let X_i be $\text{Ber}(p)$ r.v.'s, $p \in (0, 1)$. Then

$$\Pr \left[\left| \sum_{i=1}^k X_i - pk \right| \geq \epsilon n \right] \leq 2 \exp(-\Theta(\epsilon^2 k))$$

In fact particular,

$$\Pr \left[\frac{1}{k} \sum_{i=1}^k X_i - p \geq \epsilon \right] \leq \exp(-\Theta(\epsilon^2 k))$$

[Assume: $p + \epsilon \leq 1$]

Pf:- We try to use all moments at once!

$$P_v \left[\sum_{i=1}^k X_i \geq (p+\epsilon)k \right]$$

$$= P_v \left[\exp \left(\lambda \sum_{i=1}^k X_i \right) \geq \exp \left(\lambda k (p+\epsilon) \right) \right]$$

$$\forall \lambda > 0.$$

$$\leq \exp(-\lambda k t) E \left[\exp \left(\lambda \sum_{i=1}^k X_i \right) \right]$$

$$= \exp(-\lambda k t) E \left[\exp(\lambda X_1) \right]^k$$

by independence
and identical
distribution.

$$= \exp(-k(\lambda t - \log E[\exp(\lambda X_1)]))$$

$$\leq \exp(-k \varphi_X(t))$$

where

$$\varphi_X(t) = \sup_{\lambda > 0} (\lambda t - \log E[\exp(\lambda X_1)])$$

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Need to compute
this.

$E[\exp(\lambda X)]$, seen as, a function of λ ,
is similar to the 'Laplace transform'.

! Now, for the Bernoulli case.

$$E[\exp(\lambda X)] \leftarrow \cancel{E[1-X + Xe^\lambda]} \\ = 1-p + pe^\lambda$$

$$\text{So, } \psi_X(t) = \sup_{\lambda \geq 0} (\lambda t - \log(1-p + pe^\lambda))$$

$$f(\lambda) = \lambda t - \log(1-p + pe^\lambda)$$

$$f'(\lambda) = t - \frac{pe^\lambda}{1-p+pe^\lambda}$$

So, $f''(\lambda) < 0$ everywhere, so zero of $f'(\lambda)$ is a maximum. This is given by

$$t(1-p+pe^\lambda) = pe^\lambda$$

$$\text{or, } \frac{t(1-p)}{p(1-t)} = e^\lambda$$

$$\Rightarrow \lambda = \log\left(\frac{t(1-p)}{p(1-t)}\right).$$

So, we get

$$\begin{aligned} Y_X(t) &= t \log \left(\frac{t}{p} \frac{1-p}{1-t} \right) \\ &\quad - \log \left(1-p + \frac{t(1-p)}{1-t} \right) \\ &= t \log \left(\frac{t(1-p)}{p(1-t)} \right) - \log \left(\frac{1-p}{1-t} \right) \\ &= t \log \left(\frac{t}{p} \right) + (1-t) \log \left(\frac{1-t}{1-p} \right) \\ &= D(t||p). \end{aligned}$$

So, $\Pr(\sum X_i - pk > \epsilon k) \leq \exp(-k D(t||p))$
where $t = p + \epsilon$.

Now $D(p+\epsilon||p)$

$$\begin{aligned} &= (p+\epsilon) \log \left(1 + \frac{\epsilon}{p} \right) + (1-p-\epsilon) \log \left(\frac{1-p-\epsilon}{1-p} \right) \\ &= (p+\epsilon) \log \left(1 + \frac{\epsilon}{p} \right) + (1-p-\epsilon) \log \left(1 - \frac{\epsilon}{1-p} \right) \\ &\quad \left(\leq (p+\epsilon) \left(\frac{\epsilon}{p} \right) \right. \\ &\quad \quad \left. + (-p-\epsilon) \left(-\frac{\epsilon}{1-p} \right) \right) \\ &\quad \left(1 + \frac{\epsilon}{p} - \epsilon \right) \\ &\quad \left(\frac{\epsilon^2}{p} + \frac{\epsilon^2}{1-p} \right) \end{aligned}$$

$$f(\varepsilon) = (p+\varepsilon) \log(p+\varepsilon) + (1-p-\varepsilon) \log(1-p-\varepsilon)$$

$$f(0) = 0 \quad \rightarrow (p+\varepsilon) \log p - (1-p-\varepsilon) \log(1-p)$$

$$f'(\varepsilon) = \log\left(\frac{p+\varepsilon}{1-p-\varepsilon}\right) - \log\left(\frac{p}{1-p}\right)$$

$$f'(0) = 0$$

$$f''(\varepsilon) = \frac{1}{p+\varepsilon} + \frac{1}{1-p-\varepsilon} \geq 4$$

$$\Rightarrow D(p+\varepsilon \| p) \geq 2\varepsilon^2$$

\Rightarrow So we get the Chernoff bdd,
and hence the correct tail:
We have to choose k so that

$$\text{Need } \exp(-2k\varepsilon^2) \leftarrow$$

$$\exp(-2k(\frac{\delta}{5})^2) \leq \delta$$

$$\text{So, } k \geq 5 \log(1/\delta) \text{ suffices.}$$

Reducing storage - Note that we used that the rows are mutually independent, but ∇ within the each row, we only needed four-wise independence.

Thm - There is a polynomial time algorithm which takes as input $O(k \log n)$ uniform random bits (denoted as R) and an index $i \in [n]$, and outputs in ~~time~~ polynomial time ($\text{poly}(k \log n)$) a bit $A(R, i)$ s.t. $\forall i_1, i_2, i_3, \dots, i_k$ distinct and in $[n]$

$$A(R, i_1), A(R, i_2), \dots, A(R, i_k)$$

are k wise independent.

$$\Pr \left[\bigwedge_{\ell=1}^k A(R, i_\ell) = b_\ell \right] = \prod_{\ell=1}^k \Pr [A(R, i_\ell) = b_\ell]$$

$$\forall (b_1, \dots, b_k) \in \{0, 1\}^k.$$

Thus we can generate each row with only $O(\log m)$ [rather than $\Theta(m)$] stored random bits.

Consider first the case $k=2$. Let r_1, r_2, \dots, r_t be uniformly random bits. Then

Claim \rightarrow The $\{ \bigoplus_{i \in S} r_i \mid S \subseteq [t], S \neq \{\emptyset\} \}$ is a set of ~~pair~~ $2^t - 1$ pairwise independent random bits.

Pf:- $\Pr(\text{Let } r_S = \bigoplus_{i \in S} r_i)$

Then, clearly, when $S \neq \{\emptyset\}$,
 $\Pr[r_S = 0] = \frac{1}{2}$.

Now Suppose $S \neq T$ are distinct non-empty sets. Let ~~let~~ ~~why~~ suppose $\exists a \in S \setminus T$ (if not ~~repl~~ ~~make T into S~~ ~~relabel~~ ~~relabel S & T as~~ each other).

Now

$$\begin{aligned} \Pr[r_S = \alpha \wedge r_T = \beta] &= \Pr[r_S = \alpha \mid r_T = \beta] \cdot \frac{1}{2} \\ &= \frac{1}{2} \cdot \frac{1}{2} = \Pr[r_S = \alpha] \Pr[r_T = \beta] \end{aligned}$$

k-wise independence:

Consider another construction. Let \mathbb{F}_p be a large enough finite field. Let a_0, a_1, \dots, a_{k-1} be a.a.v. chosen random elements from \mathbb{F} .

Then the elements

$$\left\{ x = \sum_{i=0}^{k-1} a_i x^i \mid x \in \mathbb{F} \right\} \text{ are } k\text{-wise}$$

independent. To see this, consider the probability

$\Pr \left[\bigwedge_{i=1}^k \alpha_i = \beta x_i \right]$ where β, x_1, \dots, x_k are distinct and $\beta x_1, \dots, \beta x_k$ are arbitrary. We have.

$\Pr \left[\bigwedge_{i=1}^k \alpha_i = \beta x_i \right] = \frac{1}{p^k}$, which proves k -wise independence.

For bits one does this by taking $p=2$ and then using a finite field structure on 2^d ($d = \lceil \lg n \rceil$).