

Functional inequalities and concentration

Last lecture :- Gaussian Poincare inequality: For "every" $f: \mathbb{R}^n \rightarrow \mathbb{R}$

$$\text{Var} [f(X)] \leq E [\|\nabla f(X)\|^2]$$

when $X \sim N(0, I^n)$.

Stronger inequality :- Gaussian log-Sobolev inequality

$$\text{Ent}(f^2) \leq 2 E [\|\nabla f(X)\|^2] \quad \text{for "every" } f: \mathbb{R}^n \rightarrow \mathbb{R}$$

when $X \sim N(0, I_n)$

where

$$\text{Ent}(f) = E \left[f(X) \log \frac{f(X)}{E[f(X)]} \right] = E[f \log f] - E[f] \log E[f]$$

Herbst argument :- "Log Sobolev inequalities imply concentration".

Let $Z = f(X)$ where $f: \mathbb{R}^n \rightarrow \mathbb{R}$ is Lipschitz, $E[Z] = 0$,

The Herbst argument deduces a bound on the log moment generating fn. $\psi_Z(\lambda) := \log E[e^{\lambda Z}]$ from the Log-Sobolev inequality.

Define $F(\lambda) = E[e^{\lambda Z}]$ and apply the log-Sobolev inequality with the fn. $z \mapsto \exp^{\lambda f(z)/2}$.

Then we have

$$\text{Ent}(Z) \leq \frac{1}{2} E [\lambda^2 \exp(\lambda f(X)) \|f(X)\|^2] \leq \frac{1}{2} \lambda^2 L^2 F(\lambda)$$

$$\Leftrightarrow \lambda E[Z e^{\lambda Z}] - F(\lambda) \log F(\lambda) \leq \frac{\lambda^2 L^2}{2} F(\lambda).$$

$$\lambda F'(\lambda) - F(\lambda) \log F(\lambda) \leq \frac{\lambda^2 L^2}{2} F(\lambda) \quad (*)$$

Dividing by $\frac{1}{\lambda^2} F(\lambda) > 0$, we get,

$$G'(\lambda) \leq 2L^2$$

$$\text{where } G(\lambda) = \frac{\log F(\lambda)}{\lambda}.$$

$$(G'(\lambda) = \frac{1}{\lambda} \frac{F'(\lambda)}{F(\lambda)} - \frac{1}{\lambda^2} \log F(\lambda))$$

We now use, $\lim_{\lambda \downarrow 0} G(\lambda) = E[Z]$, so we get.

$$\log E[\exp(\lambda Z)] = F(\lambda) \leq \exp\left(\frac{\lambda^2 L^2}{2}\right) \Rightarrow Z \in \mathcal{G}(L^2)$$

$$\Rightarrow \Pr [f(x) - E[f(x)] > t] \leq \exp\left(-\frac{t^2}{2L^2}\right).$$

[Gaussian concentration]

Interestingly, such functional inequalities often have to be proved only in dimension $n=1$; the extension to higher dimensions is automatic based on sub-additivity properties we will now see.

— Instead of following this approach, we will use a slightly different approach [still based on appropriate functional inequalities in $n=1$ dimension], which still uses such sub-additivity properties.

(*) \Rightarrow Requires some measure theoretic justification, but is valid, e.g., when $E[e^{\lambda Z}]$ exists in some interval $[a, b]$, and λ is in the interior of this interval.

Sub-additivity: The Efron-Stein inequality and the sub-additivity of Entropy.

Notation: We denote

$$E_x[f(x, Y)] = \int f(x, Y) \mu(x, Y) dx = E[X|Y]$$

more robust defn.

and if X_1, \dots, X_n is a set of r.v.'s, $S \subseteq [n]$ then

$$E^{(S)}[f(X_1, \dots, X_n)] := E[f(X_1, \dots, X_n) | X_S]$$

In particular,

$$E^{(i)}[f(X_1, \dots, X_n)] = E[f(X_1, \dots, X_n) | X_{\{1, \dots, X_{i-1}, X_{i+1}, \dots, X_n\}}]$$

For any fn Φ , and random variable Z .

define $H_{\Phi}(Z) := E[\Phi(Z)] - \Phi(E[Z])$.

Examples :-

$$\text{Var}(Z) : \Phi(Z) = Z^2$$

$$\text{Ent}(Z) = E[Z \log Z] - E[Z] \log E[Z]$$

$$\Phi(Z) = Z \log Z.$$

Conditional versions:- Let $Z = f(X_1, \dots, X_n)$

$$H_{\Phi}^{(S)}(Z) = E^{(S)}[\Phi(Z)] - \Phi(E^{(S)}[Z])$$

"function" of X_S .

$$\text{Var}^{(i)}(Z) = E^{(i)}[Z^2] - E^{(i)}[Z]^2 \quad (\text{"jth" of } X_1, \dots, X_{i-1}, X_{i+1}, \dots, X_n)$$

Efron-Stein: Assume X_1, X_2, \dots, X_n are independent. $Z = f(X_1, \dots, X_n)$

$$\text{Var}(Z) \leq \sum_{i=1}^n E[\text{Var}^{(i)}[Z]]$$

Sub-additivity of entropy: Same framework as above. X_1, \dots, X_n independent

$$\text{Ent}(Z) \leq \sum_{i=1}^n E[\text{Ent}^{(i)}[Z]]$$

Convexity: H_{Φ} is a convex functional in both the above cases: if X, Y are independent then,

$$H_{\Phi}[E_X[f(X, Y)]] \leq E_X[H_{\Phi}^{(Y)}[f(X, Y)]]$$

Claim: Convexity \Rightarrow Sub-additivity

Proof: We first derive the following identity

$$H_{\Phi}[f(X_S, X_{\bar{S}})] = E[\Phi(f(X_S, X_{\bar{S}})) - \Phi(E[f(X_S, X_{\bar{S}})])]$$

$$= E[\Phi(f(X_S, X_{\bar{S}})) - \Phi(\bar{E}_{X_S}[f(X_S, X_{\bar{S}})])]$$

$$+ \Phi(\bar{E}_{X_S}[f(X_S, X_{\bar{S}})])$$

$$- \Phi(E[E_{X_S}[f(X_S, X_{\bar{S}})])]$$

$$\begin{aligned}
&= E \left[E_{X_s} \left[\Phi \left(f(X_s, X_{\bar{s}}) \right) \right] \right] - E \left[\Phi \left(E_{X_s} \left[f(X_s, X_{\bar{s}}) \right] \right) \right] \\
&\quad + E \left[\Phi \left(E_{X_s} \left[f(X_s, X_{\bar{s}}) \right] \right) \right] \\
&\quad - \Phi \left(E \left[E_{X_s} \left[f(X_s, X_{\bar{s}}) \right] \right] \right) \\
&= E_{X_{\bar{s}}} \left[H_{\Phi}^{(s)}(f) \right] + H_{\Phi} \left(E_{X_s} \left[f(X_s, X_{\bar{s}}) \right] \right)
\end{aligned}$$

Now, when $X_s, X_{\bar{s}}$ are independent, the convexity assumption implies,

$$H_{\Phi} \left(E_{X_s} \left[f(X_s, X_{\bar{s}}) \right] \right) \leq E_{X_s} \left[H_{\Phi}^{(s)}(f) \right]$$

So, we have

$$H_{\Phi}(f) \leq E \left[H_{\Phi}^{(s)}(f) \right] + E \left[H_{\Phi}^{(\bar{s})}(f) \right] \quad \text{whenever, } X_s, X_{\bar{s}} \text{ are independent.}$$

In particular, when $(X_i)_{i=1}^n$ have a product distribution, we get

$$\begin{aligned}
H_{\Phi}(f) &\leq E \left[H_{\Phi}^{(1)}(f) \right] + E \left[H_{\Phi}^{\{2, \dots, n\}}(f) \right] \\
&\leq E \left[H_{\Phi}^{(1)}(f) \right] + E \left[H_{\Phi}^{(2)}(f) \right] + E \left[H_{\Phi}^{\{3, \dots, n\}}(f) \right] \\
&\leq \sum_{i=1}^n E \left[H_{\Phi}^{(i)}(f) \right].
\end{aligned}$$

Both Var and Ent are convex, which we can prove by showing that they are suprema of linear functionals.

(1) Variance = By definition, for any r.v. [for which $E[T^2]$ is well defined]

$$0 \leq \text{Var}(Y-T) = \text{Var}(Y) + \text{Var}(T) - 2 \text{Cov}(Y, T)$$

$$\Rightarrow \text{Var}(Y) \geq 2 \text{Cov}(Y, T) - \text{Var}(T) \quad \forall T \text{ s.t. } E[T^2] \text{ exists.}$$

with equality when $T=Y$.

So,
$$\text{Var}(Y) = \sup_{T: E[T^2] < \infty} \underbrace{2 \text{Cov}(Y, T) - \text{Var}(T)}_{\text{Linear functional of } Y}$$

$\underbrace{\sup}_{\text{sup over linear function.}} E[T(2Y-T)]$

For entropy :- Let Y be any non-negative r.v. and let T be any random variable that is zero whenever Y is. (denoted $T \ll Y$)

We use the convexity of $-\log$, which implies that

$$E\left[\frac{Y}{E[Y]} \log\left(\frac{Y/E[Y]}{T/E[T]}\right)\right] \geq -\log\left(E\left[\frac{T}{E[T]}\right]\right) = 0$$

Which implies, (using $E[Y] > 0$)

$$E[Y(\log Y - \log E[Y])] \geq E[Y(\log T - \log E[T])]$$

$$\text{Ent}(Y) \quad \forall T \ll Y \text{ non-negative (with equality if } Y=T)$$

Therefore,

$$\begin{aligned} \text{Ent}(Y) &= \sup_{\substack{T \ll Y \\ T \geq 0}} E[Y(\log T - \log E[T])] \\ &= \sup_{\substack{U \text{ s.t.} \\ E[e^U] = 1}} E[UY] \end{aligned}$$

Further :- If $E[U Y] \geq \text{Ent}(Y) \quad \forall Y$, then $E[e^U] \leq 1$, for applying to $Y = e^u$, we get

$$E[U e^U] - E[e^U] \log E[e^U] \leq E[U e^U]$$

$$\Rightarrow E[e^U] \leq 1$$

Main result :- Let P ,

$$\log E_P[\exp(\lambda(Z - \lambda Z))] = \sup_{Q \ll P} [E_Q Z - E_P Z - D(Q \| P)]$$

Using this one gets

$$\Psi_{z-\lambda z}(\lambda) \leq \frac{\nu \lambda^2}{2} \quad \text{under } P$$

\Leftrightarrow

$$E_Q Z - E_P Z \leq \sqrt{2\nu D(Q||P)} \quad \forall Q \ll P.$$